ONE DOES NOT SIMPLY

REGRESS A BINARY OUTCOME APPLYING ORDINARY LEAST SQUARES

Logistic Regression

Introduction to Statistics

made on imqu

* Short Recap: Linear Regression, Interactions, Polynomials

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 - * Where to go next, Q&A.

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* An outcome variable *Y* is generated in the population as a linear combination of variables plus some chance error *\varepsilon*:

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- * This procedure recovers the population parameters *without bias* and *efficiently* under some strong assumptions about model specification and the nature of the error term.

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- * When a variable is nominal, each category will have its own coefficient, which refers to the **expected difference** in the outcome between that category and the 'reference group'.
- Standard errors represent the uncertainty of the coefficient estimate.
 P-value summarise our evidence against the null that the coefficient is zero in the population.

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* The effect of X_1 linearly depends on X_2 . As we increase X_2 by one unit, the effect of a one-unit increase of X_1 on Y goes up by β_3 .

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- * Graphically, a parabola with vertex at $-\beta_1/2\beta_2$. U-shaped if $\beta_2 > 0$, n-shaped if $\beta_2 < 0$.
- * Slope varies across values of *X*: instantaneous rate of change is $\beta_1 + 2\beta_2 X$. (The derivative, which will come back today!)

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 constant.
- * Conditional Effect Plots. Plot
 the marginal effect of *X* on *Y*across values of *Z* (moderation)
 or *X* itself (non-linearity).



Figure 2. Marginal Effects Plot: Differentiating the Effect of Democratic and Authoritarian Reforms on Satisfaction With Democracy. Estimates from Model 5 in Table 3.

Authoritarian education reform


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* Linear Probability Model (LPM): regress a 0-1 binary variable on covariates; interpret the predicted values as fractional probabilities.







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 - * Non-constant variance: Y_i is either 0 or 1, but \hat{Y}_i can be any value. So the absolute size of the error $\epsilon_i = Y_i \hat{Y}_i$ gets smaller as \hat{Y} gets closer to 0 or 1, and bigger as it gets farther. So $Var(\epsilon)$ and \hat{Y} are correlated.

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- * Only advantages of LPMs: easy-to-interpret coefficients and computationally faster than alternative. With today's software, generally **no good reason to use them** (though still pop up in econ).

Logistic Regression: Intuition



Logistic Regression 'Squiggles'



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'Squiggles' in Multiple Dimensions

Multiple Logistic Regression



Logistic Regression

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- * Raising *e* to the power of something is the inverse of taking a natural logarithm of something:

$$e^3 = 20.08554... \rightarrow \log(20.08554...) = 3$$

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* $\log \frac{\Pr(Y=1)}{1 - \Pr(Y=1)}$ is known as **log-odds**, or **logit** function of $\Pr(Y=1)$.



Log-Odds

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Log-odds centre the outcome at 0 and linearise it:

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- * For p = 0.9, odds is $9 \rightarrow \text{log-odds or logit}(0.9)$ is $\log(9) \approx 2.20$
- * For p = 0.1, odds is $1/9 \rightarrow \text{log-odds or logit}(0.1)$ is $\log(1/9) \approx -2.20$





log-odds







log-odds





probability (0-1)















0.75







Х



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- * logit(0.75) ≈ 1.10

- * $logit^{-1}(-1.10) \approx 0.25$
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- * logit(0.25) ≈ -1.10
- * logit(0.5) = 0
- * logit(0.75) ≈ 1.10
- * $logit(0.952) \approx 3$

- * $logit^{-1}(-1.10) \approx 0.25$
- * $logit^{-1}(0) = 0.5$
- * $logit^{-1}(1.10) \approx 0.75$
- * $logit^{-1}(3) \approx 0.952$

Logistic Regression, Two Ways

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With *Y* as a probability:

 $Pr(Leave = 1) = \frac{1}{1 + e^{-(\alpha + \beta Euroscepticism)}}$

 $Pr(Leave = 1) = logit^{-1}(\alpha + \beta Euroscepticism)$

- * Easy-to-interpret left-hand side: it's a probability, can only take values comprised between 0 and 1.
- * Hard-to-interpret right-hand side: it's a nonlinear curve (sigmoid). Not obvious what a 1-unit increase in *X* does.

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With *Y* as log-odds:

 $\log \frac{\Pr(\text{Leave} = 1)}{\Pr(\text{Leave} = 0)} = \alpha + \beta \text{Euroscepticism}$

 $logit[Pr(Leave = 1)] = \alpha + \beta Euroscepticism$

- * Easy-to-interpret right-hand side: it's a linear function, like with the linear model. A 1-unit increase in X increases outcome by β.
- Hard-to-interpret left-hand side: it's a funky way of expressing probabilities, which can take any value from – inf to + inf.

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	Dependent variable:
	Leave Vote
Intercept	-3.68 (2.63)
Euroscepticism	0.56 (0.38)

Observations 7



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- * Log-odds of Leave vote when Euroscepticism = 1: -3.68 + 0.56 = -3.12

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Dan and dant up in allo



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- ★ Log-odds of Leave vote when Euroscepticism = 0: -3.68
- * Log-odds of Leave vote when Euroscepticism = 1: -3.68 + 0.56 = -3.12
- * Log-odds of Leave vote when Euroscepticism = 2: $-3.68 + 2 \times 0.56 = -2.56$

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- * Probability of Leave vote for Euroscepticism = 1 $logit^{-1}(-3.68 + 0.56) = 0.042$

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- * Probability of Leave vote for Euroscepticism = 0 $logit^{-1}(-3.68) = 0.024$
- * Probability of Leave vote for Euroscepticism = 1 $logit^{-1}(-3.68 + 0.56) = 0.042$
- * Probability of Leave vote for Euroscepticism = 2 $logit^{-1}(-3.68 + 2 \times 0.56) = 0.072$

	Dependent variable:
	Leave Vote
Intercept	-3.68 (2.63)
Euroscepticism	0.56 (0.38)



- * Use **inverse-logit function** to get the predicted probability:
- * Probability of Leave vote for Euroscepticism = 0 $logit^{-1}(-3.68) = 0.024$
- * Probability of Leave vote for Euroscepticism = 1 $logit^{-1}(-3.68 + 0.56) = 0.042$
- * Probability of Leave vote for Euroscepticism = 2 $logit^{-1}(-3.68 + 2 \times 0.56) = 0.072$
- * Probability of Leave vote for Euroscepticism = 3 $logit^{-1}(-3.68 + 3 \times 0.56) = 0.12$

	Dependent variable:
	Leave Vote
Intercept	-3.68 (2.63)
Euroscepticism	0.56 (0.38)



	Dependent
	Leave Vote
Intercept	-6.21 (0.93)
Euroscepticism	0.78 (0.13)
Johnson Approval	0.26 (0.09)
Observations	200



With multiple predictors, the
 change in log-odds associated
 with each predictor is still linear.

	Dependent
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 Euroscepticism depends both on the level of Euroscepticism and on the level of Johnson
 Approval...
- In complex models, interpret sign and significance of coefficients, do not interpret their value.

	Dependent
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Logistic Regression in R

```
> model <- glm(Leave ~ Euroscepticism + likeJohnson, data = bes,
+ family = "binomial")
> model <- glm(Leave ~ Euroscepticism + likeJohnson, data = bes,
+ family = binomial(link = "logit"))
> model <- glm(Euroscepticism ~ likeJohnson, data = bes,
+ family = "binomial")
Error in eval(family$initialize) : y values must be 0 <= y <= 1
>
```

 ∗ Use when the dependent variable is a 0 − 1 binary variable, and we want to know the **probability** that it takes the value of 1.

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- But you can safely interpret sign (±) and significance: "Euroscepticism is positively and significantly (*p* < 0.05) associated with probability of Leave vote, holding all else constant."
- * Use **predicted values plot** to get a sense of substantive effects for an 'average' observation, expressed in terms of predicted probabilities.

Logistic Regression: Two Extra Steps

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* Average Marginal Effects

- * Summarise average relationship between the regressors and the outcome in terms of probability.
- * Useful quantity to **interpret** of model estimates, and increasingly common. But not integral or specific to logistic regression.

Logistic Regression: Two Extra Steps

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* Maximum Likelihood Estimation

- How your statistical software picks a particular set of coefficients (i.e. a particular 'squiggle') over all possible others.
- * Essential to the **computation** of model estimates. But R does it for you, so it's just nice to have a vague idea of what's going on.
For an observation *i*, we can get the marginal effect, or the instantaneous rate of change in probability at one point with the derivative at its predicted value.

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X	Marginal Effects	1 ===
1	0.023	
3	0.058	



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1	0.023
3	0.058
4	0.086



V	Monsing 1 Effects	1	
Λ	Marginal Effects		
1	0.023		
3	0.058		
4	0.086	ave)	
5	0.115	p(le	
		$\frac{dy}{dx}$	= 0.115



		1	
X	Marginal Effects		
1	0.023		$\frac{dy}{dx} = 0.120$
3	0.058		
4	0.086	ave)	
5	0.115	p(le	
8	0.120		

		1	_
X	Marginal Effects		
			$\frac{dy}{dx} = 0.091$
1	0.023		
3	0.058		
4	0.086	eave)	
5	0.115	p(le	
8	0.120		
9	0.091		

5 6 7

Euroscepticism

		1 —	
Х	Marginal Effects		
			$\frac{dy}{dx} = 0.063$
1	0.023		
3	0.058		
4	0.086	eave)	
5	0.115	p(le	
8	0.120		
9	0.091		
10	0.063		
		•	
		0 -•	

Euroscepticism

		1
X	Marginal Effects	
1	0.023	
3	0.058	
4	0.086	Save)
5	0.115	e e e e e e e e e e e e e e e e e e e
8	0.120	
9	0.091	
10	0.063	
Mean	0.080	

Euroscepticism

 $Pr(Leave) = logit^{-1}(\alpha + \beta_1 Euroscepticism + \beta_2 Trust + \beta_3 Gender + \epsilon)$

R# Euroscepticism Trust in MPs Female M.E. (Eurosc.) M.E. (Trust in MPs) ME (Female)

R#	Euroscepticism	Trust in MPs	Female	M.E. (Eurosc.)	M.E. (Trust in MPs)	ME (Female)
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2	10	2	0	0.021	-0.007	0.003
3	2	4	1	0.008	-0.003	0.001

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5	10	4	0	0.041	-0.014	0.006
6	0	3	1	0.001	-0.001	0.0002
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8	10	4	0	0.041	-0.014	0.006
Average Marginal Effects (=Mean)				0.078	-0.027	0.013

<pre>> margins(model)</pre>
Average marginal effects
<pre>glm(formula = Leave ~ Euroscepticism + trustMPs + gender, family = "binomial", data = bes</pre>
Euroscepticism trustMPs gender
0.07841 -0.02719 0.01259
<pre>> margins_summary(model)</pre>
factor AME SE z p lower upper
Euroscepticism 0.0784 0.0014 55.9855 0.0000 0.0757 0.0812
gender 0.0126 0.0394 0.3195 0.7493 -0.0646 0.0898
trustMPs -0.0272 0.0128 -2.1195 0.0340 -0.0523 -0.0020
<pre>> head(marginal_effects(model))</pre>
dydx_Euroscepticism dydx_trustMPs dydx_gender
1 0.034934821 -0.0140799805 0.0056085222
2 0.021075526 -0.0086488138 0.0033835164
3 0.007761427 -0.0032417654 0.0012460383
4 0.048529412 -0.0192219974 0.0077910323
5 0.040877210 -0.0163496473 0.0065625284
6 0.001454886 -0.0006128382 0.0002335708

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- * Why 'almost' correct? Because AMEs aren't averages of one-point changes (these would be slopes that go from \hat{Y}_i for X_i and the value of \hat{Y}_i for $X_i + 1$).

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- * AMEs are averages of slopes **at one point** (i.e. derivatives).
- * But the 'one-point increase' interpretation is fine. Or just say: the **average marginal effect** of Euroscepticism is 7.8 percentage points.


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Fig. 2 Average marginal effects on IMI-support

Note: plot displays the average marginal effects with 95% confidence intervals for the independent variables on IMI support. Estimates based on a logistic regression model with standardized independent variables (for detailed model output see Model 1 in Table A-2 in the supplementary materials)

data from oursangpleanaot.from2hypothetical 'average cases'.

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Figure 4. Average marginal effects of class, union membership, attitudes on the different support groups of social democracy. Note: The figure shows average marginal effects based on the models presented in Table B.2. The reference category for social class is socio-cultural professionals. The left-hand side shows the contrast between demobilised and core supporters, whereas the right-hand side shows the contrast between demobilised and distant supporters is shown in Supplementary Appendix D.2.

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Figure 4. Average margi figure shows average mar cultural professionals. Th shows the contrast betw Supplementary Appendix DEST APPTO2 from limite



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Discrete changes from 'improved' for 'stayed same' and 'declined'. Discrete changes from 'will improve' for 'will be same' and 'will decline'.

Figure 6. Average marginal effects of experience and expectation of status decline on party choice based on Model 2 and using the 2018 survey data set.

Notes: The regression model includes sociodemographic controls and respondents' left-right and liberal-conservative political ideologies (see Table A5 in supplementary material). Whiskers represent 95% confidence intervals. When they intersect the red dotted line, the difference in group means is not statistically significant (P < 0.05).

Im, Wass, Kantola and Kauppinen (2022)

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Figure 1. How social, economic, and climate programs shape support for bundled climate policy. The left panel shows average effects of each policy element (colored by policy dimension) on support for the policy bundle, while the right panel shows party-specific effects (red = Republican, blue = Democrat). Policy dimensions include carbon taxes, social programs, economic programs, energy costs, government spending levels, and party sponsorship. Point estimates are average marginal component effects (AMCEs) with 95% confidence intervals for each policy level. Each AMCE estimates how inclusion of the listed program affects support for the bundled climate package. Each element is compared against a base category for each policy dimension, denoted by an open circle.

Bergquist, Mildenberger and Stokes (2020)

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* When Leave_i = 0 we get $\log \frac{0}{1} = -\infty$ (negative infinity). So all residuals will be $-\infty$ - something = $-\infty$.

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* When Leave_i = 0 we get $\log \frac{0}{1} = -\infty$ (negative infinity). So all residuals will be $-\infty$ - something = $-\infty$.

* When Leave_i = 1 we get $\log \frac{1}{0} = +\infty$ (positive infinity). So all residuals will be $+\infty$ - something = $+\infty$.



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- * Likelihood: Pr(Data | Model).
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$$\ell(\alpha, \beta) = \prod_{Y_i=1}(\hat{p}_i) \prod_{Y_i=0}(1 - \hat{p}_i)$$





				1 —		#1					
R#	X	Y	Likelihood			πı	ł				
1	5	1	0.29				1				
							1			/	/
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				-eave					/		
				Pr(I				/			
						0.29					
						,					
				0						•	
				1	2 3	4	5	6	7	8	9

10

Euroscepticism

R#	X	Y	Likelihood
1	5	1	0.29
2	9	1	0.79



R#	X	Y	Likelihood
1	5	1	0.29
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3	10	1	0.87
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1	5	1	0.29
2	9	1	0.79
3	10	1	0.87
4	1	0	1-0.04 = 0.96
5	3	0	1-0.12 = 0.88



R	#	X	Y	Likelihood	1 =
]	l	5	1	0.29	
2	2	9	1	0.79	
C	3	10	1	0.87	
Ĺ	1	1	0	1-0.04 = 0.96	Leave
Ę	5	3	0	1-0.12 = 0.88	Pr(
(5	4	0	1-0.19 = 0.81	



R	\$ #	X	Y	Likelihood	1
-	1	5	1	0.29	
1	2	9	1	0.79	
~	3	10	1	0.87	
2	4	1	0	1-0.04 = 0.96	Leave
	5	3	0	1-0.12 = 0.88	Pr(
(6	4	0	1-0.19 = 0.81	
,	7	8	0	1-0.69 = 0.31	
					0.04



R#	X	Y	Likelihood	1#1#2	#3
1	5	1	0.29	0.79	0.87
2	9	1	0.79	0.69	
3	10	1	0.87		
4	1	0	1-0.04 = 0.96	Leave	
5	3	0	1-0.12 = 0.88	Ě	
6	4	0	1-0.19 = 0.81	0.29	
7	8	0	1-0.69 = 0.31	0.19	
Likelihood of the model 0.043			0.043	0.04	
	0.29 × × 0	< 0.79 × .88 × 0.8	0.87×0.96 81×0.31	1 2 3 4 5 6 7 8 9 Euroscepticism	10
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 Neither is particularly useful.

<pre>> stargazer(model)</pre>	, type = "text", single.row = TRU	E
	Dependent variable:	
	Leave	
Euroscepticism trustMPs genderFemale Constant	1.020*** (0.143) -0.353** (0.172) 0.164 (0.513) -5.655*** (1.031)	
Observations Log Likelihood Akaike Inf. Crit.	199 -50.830 109.660	
Note: >	*p<0.1; **p<0.05; ***p<0.01	

- * No R^2 , for the same reason that we can't do OLS: squared residuals are all ∞ .
- * Software normally returns the Log-Likelihood (LL), and something called the Akaike Information Criterion.
 Neither is particularly useful.
- * Pseudo-R² is a bit better:
 compares the LL of the model
 with the LL of a model
 without independent
 variables.

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Our Model











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- * Where *p* is the number of independent variables.
- * A number of other pseudo- R^2 (Cox&Snell, Nagelkerke, Tjur).
- * Any of these goodness-of-fit statistics may be useful for model comparison, but **not worth losing your sleep over**.

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- * Direction and significance of log-odds coefficients are easily interpretable; effect size is not.
- * Average Marginal Effect is the best way to express how change in *X* affects the probability that *Y* is 1.
- * Use predicted probabilities and AMEs to get a sense of the substantive relationships between variables.



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 - * Other non-linear models (multinomial, Poisson \rightarrow extensions of logistic regression).
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- * Method options are sprawling and changing fast (AI is coming for all of us) make your methods training fit your research needs, not the other way around.


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 - * Causal Inference for a taste, see Imbens (forthcoming) "Causal Inference in the Social Sciences", Annual Review of Statistics and Its Application.
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- * Long-term investment will involve some self-learning.

Thank you for your kind attention!

Leonardo Carella leonardo.carella@nuffield.ox.ac.uk